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24 INTRODUCTION

25 Inter-limb asymmetries refers to the concept of the performance of two limbs not being equal
26 (3,21) and have been a popular source of investigation in recent years. Historically, many
27 studies have highlighted the prevalence of inter-limb asymmetry across a range of tests such
28 as the back squat (1,24,33), isometric squats or mid-thigh pulls (10,14,15), and jumping-based
29 tasks (2,20,28,31). Although interesting, their prevalence alone does little to enhance our
30 understanding of whether these differences should be corrected during training. More recently,
31 studies have aimed to investigate whether such asymmetries are detrimental to physical or
32 sports performance (6) with equivocal findings. For example, Hart et al. (15) showed that
33 asymmetries in strength of ~8% were associated with reduced kicking accuracy, whilst
34 Dos'Santos et al. (10) reported no association between strength asymmetries (~13%) and
35 performance during the 505 change of direction speed test. Similarly, Dos'Santos et al. (11)
36 reported no association between single and triple leg hop asymmetries and change of direction
37 speed (CODS) performance, although it should be noted that the reported inter-limb differences
38 of ~7% can be considered small (6). In contrast, Bishop et al. (4) showed that both vertical and
39 horizontal asymmetries were associated with reduced jump ($r = -0.47$ to -0.56) and sprint ($r =$
40 0.49 to 0.59) performance in elite youth female soccer players. Consequently, this lack of
41 agreement highlights the need for further research.

42 The majority of literature relating asymmetries to physical performance measures have used
43 jump tests to quantify the asymmetry component (4,11,18,26,27). Inter-limb differences from
44 horizontal jumping (such as single, triple, and crossover hop tests) have reported asymmetries
45 of 6-7% (4,11,30). When vertical asymmetries have been assessed via a single leg
46 countermovement jump (SLCMJ), these differences have been shown to be significantly
47 greater than horizontal tests (4,26,29), with values $> 10\%$ common for this test. Finally, the use
48 of drop jumps has highlighted individual asymmetry values $> 50\%$ (28) in healthy adult

49 populations; thus, the available body of evidence would suggest that the magnitude of
50 asymmetries are test-specific.

51 In addition to this varying magnitude, recent studies have displayed individual athlete
52 asymmetry data highlighting that both the left and right (4,27) or dominant and non-dominant
53 limbs (11,13) have the potential to score higher during jump testing. Despite these recent
54 findings, to the authors' knowledge, no studies to date have used this approach to specifically
55 examine if the levels of agreement in asymmetry (right versus left) is consistent across multiple
56 tests. For example, if peak force (PF) data was obtained during two different types of unilateral
57 jumps, such as a SLCMJ and single leg broad jump (SLBJ); would the same limb always record
58 the larger peak force value despite the tests being different. Therefore, the aim of the present
59 study was to assess if inter-limb asymmetries consistently occurred on the same limb during
60 unilateral strength and power tests. When reporting inter-limb differences, it was hypothesised
61 that both the magnitude and side which favoured the asymmetry would be test and metric-
62 specific, and highly individual in nature, justifying the need for an individual approach to data
63 analysis.

64

65 **METHODS**

66 *Experimental Approach to the Problem*

67 The present study required subjects to partake in two sessions. The first visit was for test
68 familiarisation. Subjects were provided with the relevant test instructions and the opportunity
69 to practice each assessment until they reached a satisfactory level of technical competence
70 during each test (established by an accredited strength and conditioning coach). Data collection
71 took place on the second visit. Subjects performed three trials on each limb for the following
72 tests: unilateral isometric squats, SLCMJ and SLBJ on a single force platform (PASPORT

73 force plate, PASCO Scientific, California, USA) sampling at 1000 Hz. Test order was
74 randomized so as to negate any potential learning effects.

75

76 *Subjects*

77 Twenty-eight recreational sport athletes (age = 27.29 ± 4.6 years; mass = 80.72 ± 9.26 kg;
78 height = 1.81 ± 0.06 m) volunteered to take part in this study. A minimum of 27 participants
79 was determined from a priori power analysis using G*Power (Version 3.1, University of
80 Dusseldorf, Germany) implementing statistical power of 0.8 and a type 1 alpha level of 0.05
81 which has been used in comparable literature (10). Inclusion criteria required all participants
82 to have a minimum of one year of resistance training experience. In addition, participants were
83 excluded from the study if they had a history of lower body injury or were injured at the time
84 of testing. Participants were required to complete informed consent forms to demonstrate that
85 they were willing and able to undertake all testing protocols. Ethical approval was granted from
86 the Research and Ethics Committee at the London Sport Institute, Middlesex University.

87

88 *Procedures*

89 A standardised dynamic warm up was conducted prior to each session consisting of dynamic
90 stretches to the lower body (such as multi-planar lunges, inchworms, and ‘world’s greatest
91 stretch’), in addition to three practice trials at 60, 80, and 100% perceived effort. Two minutes
92 of rest was provided after the final warm up trial before undertaking the first test. It should be
93 noted that although additional metrics could be quantified from the force platform, only
94 comparable metrics across tests were computed given the focus of this study was to establish
95 asymmetry side consistency across the different tests. Finally, although test order was
96 randomised, trials were always conducted on the left limb first.

98 *Unilateral Isometric Squat.* A custom built 'ISO rig' (Absolute Performance, Cardiff, UK) was
99 used for this test protocol. A goniometer was used to measure $\sim 140^\circ$ of hip and knee flexion
100 (14) for each participant, with full extension of the knee joint equalling 180° . The fulcrum of
101 the goniometer was positioned on the lateral condyle of the femur. The stabilisation arm was
102 lined up along the line of the fibula (in the direction of the lateral malleolus) and the movement
103 arm was lined up with the femur (pointing towards the greater trochanter at the hip). The non-
104 stance limb was required to hover next to the working limb, so as to try and keep the hips level
105 during the isometric squat action; thus, aiding balance and stability. Once in position,
106 participants were required to remain motionless for 2-seconds, without applying any upwards
107 force (which was verified by manual detection of the force-time curve in real time). Each trial
108 was then initiated by a "3, 2, 1, Go" countdown and participants were instructed to try and
109 extend their knees and hips by driving up as "fast and hard as possible" (10) against the bar for
110 three seconds. PF was recorded and was defined as the maximum force generated during the
111 test and reported as absolute values.

112

113 *Unilateral Countermovement Jump.* Participants were instructed to step onto the force plate
114 with their designated test leg with hands placed on hips which were required to remain in the
115 same position for the duration of the test. The jump was initiated by performing a
116 countermovement to a self-selected depth before accelerating vertically as explosively as
117 possible into the air (34). The test leg was required to remain fully extended throughout the
118 flight phase of the jump before landing back onto the force plate as per the set up. The non-test
119 leg was flexed at the hip to $\sim 90^\circ$ for the duration of each trial. Each trial was separated by 60
120 seconds of rest. Recorded metrics for each trial included PF (propulsive), eccentric and

121 concentric impulse, with definitions for their quantification conducted in line with suggestions
122 by Chavda et al. (7). Peak propulsive force was defined as the maximum force output during
123 the propulsive phase of the jump. Eccentric impulse was defined as the force exerted multiplied
124 by the time taken to produce it during the eccentric braking phase of the jump. Concentric
125 impulse was defined as the force multiplied by the time taken to produce it during the
126 concentric propulsion phase of the jump (7).

127

128 *Unilateral Broad Jump.* Participants stood on the force plate with their designated test leg and
129 hands placed on their hips. The jump was initiated by performing a countermovement to a self-
130 selected depth before jumping forward as far as possible (34). The fronts of the participants'
131 shoes were placed on the edge of the force plate (without going over) so that the edge of the
132 force plate also served as 0 cm. The tape measure (which was fixed to the floor) ran
133 perpendicular to the force plate for distance to be measured from the heel of the landing foot.
134 Participants were required to “stick the landing” and avoid toppling forward, otherwise trials
135 were excluded and subsequently retaken after a 60-second rest interval. Recorded metrics
136 included PF, eccentric and concentric impulse respectively.

137

138 *Statistical Analyses*

139 Initially all force-time data were exported to Microsoft Excel™, expressed as means and
140 standard deviations (SD), and later transferred into SPSS (V.24, Chicago, IL, USA) for
141 additional analyses. Within-session reliability was quantified for each metric in both test
142 sessions using the coefficient of variation (CV: $SD[\text{trials 1-3}]/\text{average}[\text{trials 1-3}]*100$) and
143 intraclass correlation coefficient (ICC) with absolute agreement. CV values < 10% were
144 deemed acceptable (9) and ICC values were interpreted in line with suggestions by Koo and

145 Li, (22) where scores > 0.9 = excellent, 0.75-0.9 = good, 0.5-0.75 = moderate, and < 0.5 = poor.
146 Noting that asymmetries may favour either the left or right limbs, a Kappa coefficient was
147 calculated to determine the levels of agreement between asymmetries for a common metric
148 across two tests (8). This method was chosen because the Kappa coefficient describes the
149 proportion of agreement between two methods after any agreement by chance has been
150 removed (8). In addition, only metrics that were common across more than one test were used
151 for this statistic (e.g., PF for all tests). Intuitively, this made sense given that asymmetries have
152 been shown to be both task and metric-specific (4,26,27,28,29). Kappa values were interpreted
153 in line with suggestions from Viera and Garrett (35), where 0.01-0.20 = slight, 0.21-0.40 = fair,
154 0.41-0.60 = moderate, 0.61-0.80 = substantial, and 0.81-0.99 = almost perfect. Finally, inter-
155 limb asymmetries were quantified as a percentage difference between limbs (from best trials)
156 using the formula proposed by Bishop et al. (4). Given that the quantification of asymmetry
157 was focused on percentage difference between limbs, no reference value was required (4). In
158 addition, it has been suggested that the easiest way to utilise this formula is in Microsoft
159 Excel™ (4); thus, a modification was made via the use of an 'IF function' (Equation).
160 Consequently, if an asymmetry score was positive, the right limb had the largest score between
161 limbs and vice versa for a negative asymmetry outcome (19).

162

163 *Equation: ((100/(maximum value))*(minimum value)*-1+100)*IF(left<right,1,-1)*

164

165 **RESULTS**

166 Mean values, asymmetries, and reliability data are presented in Table 1. Results showed
167 moderate to excellent reliability (ICC) and acceptable consistency (CV) for each test and
168 metric. Levels of agreement for inter-limb asymmetry scores were calculated using the Kappa

169 coefficient and are shown and described in Table 2. Results showed slight to fair levels of
170 agreement (range = -0.34 to 0.32) for all comparisons with the exception of concentric impulse
171 between the SLCMJ and SLBJ (0.79) which showed substantial levels of agreement. Individual
172 asymmetry values for PF (across all tests) are shown in Figure 1, and for eccentric and
173 concentric impulse for the SLCMJ and SLBJ in Figure 2. It has been suggested that
174 asymmetries may only be 'real' if greater than the test variability (3,5,12), which in this study
175 is represented by the CV value. Thus, the reader is encouraged to pay particular attention to
176 Figures 1 and 2 where the asymmetry bars surpass the dotted line (which represents the largest
177 CV value for those given metrics).

178

179 *** INSERT TABLES 1-2 ABOUT HERE ***

180 *** INSERT FIGURES 1-2 ABOUT HERE ***

181

182 **DISCUSSION**

183 The aim of the present study was to show whether inter-limb asymmetries were favoured for
184 the same limb during the unilateral isometric squat, SLCMJ and SLBJ tests. Test reliability
185 was generally good to excellent; however, levels of agreement for measures of peak force,
186 eccentric and concentric impulse across tests was poor, with the exception of concentric
187 impulse between the SLCMJ and SLBJ. This was also represented by individual asymmetry
188 analyses (Figures 1-2). These data indicate that asymmetries are test-specific, highly individual
189 in nature, and rarely favour the same limb when comparing across tests.

190 Mean scores, mean asymmetry, and reliability data are presented in Table 1. When asymmetry
191 values are considered, previous research has suggested that ~10% might be considered a
192 potential threshold where reductions in performance (6) and heightened injury risk occur

193 (23,32). Therefore, mean asymmetry values in the present study can be considered small. Of
194 note though (and although these mean values are small), it is interesting to see all jumping-
195 based asymmetry values favour the left limb (as represented by negative scores) and the
196 isometric squat favouring the right limb (positive asymmetry outcome). Thus, the asymmetry
197 values alone highlight how one limb may be favoured over the other from task to task. Although
198 somewhat anecdotal, it is plausible that the majority of subjects' right limb was their dominant
199 (often defined by the preferred kicking limb) (15,16,17), which has been shown to be
200 outperformed by the non-dominant limb in previous research (13,15). However, due to the wide
201 range of sporting experience from the present sample and the calculation of asymmetry focused
202 on a percentage difference at a given time point, no reference value (i.e., dominant vs. non-
203 dominant) was defined.

204 Table 2 shows the results of the Kappa agreement between metric analysis. The Kappa
205 coefficient describes the proportion of agreement between two methods after any agreement
206 by chance has been removed (8). In the present study, PF was a common metric across all tests;
207 thus, asymmetry values were comparable (Figure 1). Noting that the present study aimed to
208 determine how common it was for asymmetries to be present on the same limb, the Kappa
209 values highlight slight to fair levels of agreement for PF asymmetries. For example, if an
210 asymmetry was favoured on the right limb during the SLCMJ, it was likely that the right limb
211 was not favoured during the isometric squat (Kappa = 0.04) or SLBJ (Kappa = 0.05),
212 remembering that this statistic removes the possibility that this agreement may have occurred
213 by chance. Where jumps were concerned, eccentric and concentric impulse metrics were
214 comparable; thus, asymmetry scores for these metrics were compared (Figure 2). Of note, a
215 comparison between concentric impulse across both jumps showed substantial levels of
216 agreement, indicating that asymmetries for this metric were often present on the same side.
217 This may indicate that a similar strategy was adopted prior to take off regardless of whether

218 the focus was maximal jump height or distance. As a result, asymmetries often appear to be
219 affected in the same way for this metric during vertical and horizontal jumping tasks. When all
220 other comparisons were drawn for impulse asymmetries, slight to fair levels of agreement were
221 present, again highlighting the individual nature of asymmetries across tests. This is in
222 agreement with previous research (4,29), although to the authors' knowledge, levels of
223 agreement in respect to asymmetries are limited to date (24,25). These results demonstrate the
224 changing nature of asymmetries from test to test, and highlights the need for a more individual
225 approach to data analysis.

226 Individual asymmetry data for PF and eccentric and concentric impulse measures are shown in
227 Figures 1 and 2 respectively. The largest mean asymmetry value for any test was -5.3%;
228 however, it is clear from both figures that many individual asymmetry values greatly surpassed
229 this. It was not uncommon for asymmetries to be > 10% across all tests with some individuals
230 demonstrating values between 20-30%. If proposed thresholds of ~10% are to be accepted as
231 cut-offs where reduced performance (6) and increased risk of injury are present (23,32), then
232 Figures 1 and 2 also clearly show that many individuals may require training interventions to
233 minimise these differences. In addition, previous literature has suggested that asymmetries
234 should be reported within the context of test variability (CV) so as to determine whether the
235 between-limb difference is outside the associated error of the test (3,4,12). Noting that multiple
236 CV scores exist, the authors chose to represent the greatest CV value for each metric as a
237 proposed threshold (as represented by the dotted lines on Figures 1 and 2) to identify when
238 inter-limb differences fell outside this value. When this is considered, it is again clear that many
239 individuals showed substantial asymmetries in PF (Figure 1) and impulse metrics (Figure 2) as
240 represented by bars surpassing the dotted lines. If mean asymmetry values were interpreted
241 alone, this would not depict the full story of how imbalanced some individuals are; thus,

242 individual data analyses for side-to-side differences appears critical to further our
243 understanding of this concept.

244 Despite the aforementioned results, readers should be mindful of a couple of limitations.
245 Firstly, the present study used recreational sport athletes; thus, these findings cannot be
246 attributed to elite athlete populations. Furthermore, the very premise of this paper highlights
247 that asymmetries are both task and metric-specific, suggesting that interpreting mean data is
248 somewhat limited. Secondly, this study used a force platform to gather data relating to
249 asymmetries. Although this is not a limitation, it is worth acknowledging that not all
250 practitioners will have access to this equipment. Therefore, an alternative strategy to determine
251 whether asymmetries are favoured for the same limb is required for practitioners governed by
252 smaller budgets. As such, previous work from Bishop et al. (4) used the SLCMJ, single leg
253 hop, triple hop, and crossover hop for distance tests to show the changing nature of asymmetries
254 between tasks. Practitioners who cannot access force platforms could consider such tests to
255 determine whether asymmetries are favoured for the same side during outcome measures such
256 as jump height and distance.

257 In summary, the results of the present study show that the levels of agreement for asymmetries
258 being present on the same side are quite low and highlights the changing nature of inter-limb
259 differences across tests. In addition, individual asymmetry scores were vastly different to mean
260 values for all metrics and highlights the necessity for a more individualised approach to
261 asymmetry analysis and will likely provide a more complete picture of the presence of inter-
262 limb differences.

263

264 **PRACTICAL APPLICATIONS**

265 The findings from the present study highlight that asymmetries vary across commonly used
266 strength and jumping-based tests and that the same side is also rarely favoured. As such,
267 practitioners should always consider the individual nature of asymmetries when interpreting
268 data relative to these side-to-side differences. If the mean values alone were used for
269 interpretation, it would suggest that no action would be needed to correct the existing between-
270 limb differences. However, individual asymmetry scores were vastly different and this type of
271 analysis may offer practitioners the chance to implement training interventions to reduce these
272 side-to-side differences on a more individual level. Noting that individualized training
273 programmes can be a challenge when working with large groups of athletes (i.e., in a team-
274 sport environment), assessing individual athlete data in respect to asymmetries offers
275 practitioners a viable method of establishing which athletes may require additional exercises
276 to their existing training programme, in an attempt to optimise physical performance and
277 reduce the risk of future injury.

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373

374

375 Table 1: Mean performance data \pm standard deviations (SD), asymmetry data, and reliability
 376 data for isometric squat, countermovement, and broad jump metrics.

Test/Metric	Mean \pm SD	Mean Asymmetry (%)	CV (%)	ICC (95% Confidence Intervals)
Iso PF (L)	1597 \pm 438 N	0.8	5.4	0.94 (0.88-0.97)
Iso PF (R)	1595 \pm 397 N		5.7	0.93 (0.87-0.96)
SLCMJ PF (L)	863 \pm 204 N	-3.4	5.8	0.89 (0.80-0.94)
SLCMJ PF (R)	831 \pm 182 N		5.3	0.93 (0.87-0.96)
SLCMJ EI (L)	70 \pm 17 N·s	-4.2	8.7	0.89 (0.81-0.95)
SLCMJ EI (R)	67 \pm 17 N·s		9.1	0.83 (0.71-0.91)
SLCMJ CI (L)	152 \pm 21 N·s	-1.6	3.3	0.92 (0.86-0.96)
SLCMJ CI (R)	150 \pm 20 N·s		4.1	0.81 (0.69-0.90)
SLBJ PF (L)	732 \pm 156 N	-1.4	8.7	0.75 (0.59-0.86)
SLBJ PF (R)	722 \pm 159 N		9.3	0.80 (0.66-0.89)
SLBJ EI (L)	59 \pm 19 N·s	-5.3	11.9	0.85 (0.74-0.92)
SLBJ EI (R)	56 \pm 17 N·s		11.1	0.87 (0.77-0.93)
SLBJ CI (L)	104 \pm 17 N·s	-1.4	7.3	0.69 (0.51-0.83)
SLBJ CI (R)	102 \pm 14 N·s		8.8	0.66 (0.47-0.81)

CV = coefficient of variation, ICC = intraclass correlation coefficient, Iso = isometric, SLCMJ = single leg countermovement jump, SLBJ = single leg broad jump, PF = peak force, EI = eccentric impulse, CI = concentric impulse, L = left, R = right, N = newtons, N·s = newton seconds.

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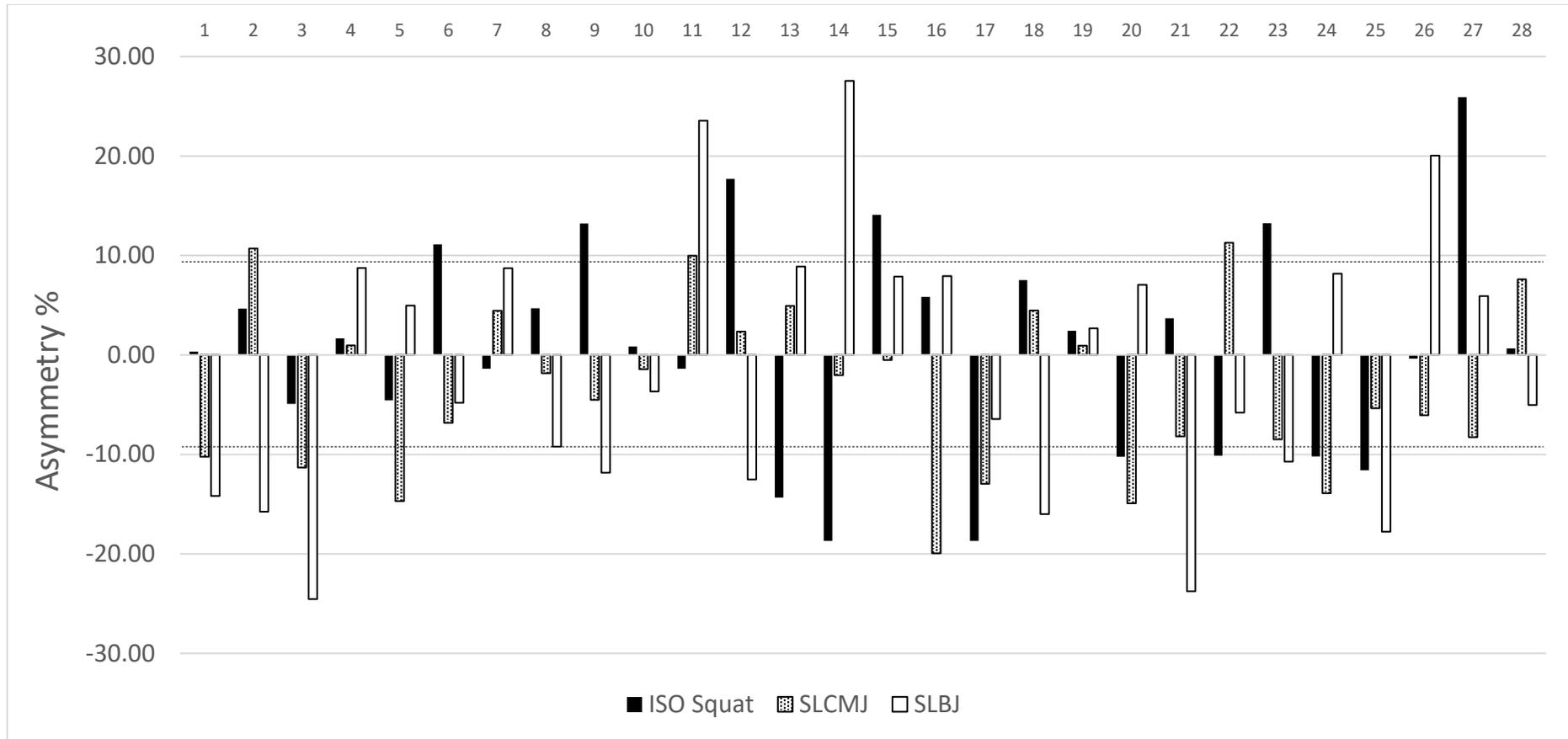
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382 Table 2: Kappa values and descriptive levels of agreement between the favored and non-
 383 favored sides for peak force, and eccentric and concentric impulse metrics across common
 384 tests.

Test Methods	Kappa Coefficient	Level of Agreement
<i>Peak Force:</i>		
Iso Squat – SLCMJ	0.04	Slight
Iso Squat – SLBJ	-0.34	Fair
SLCMJ – SLBJ	0.05	Slight
<i>Impulse:</i>		
SLCMJ Ecc – SLBJ Ecc	0.32	Fair
SLCMJ Con – SLBJ Con	0.79	Substantial
SLCMJ Ecc – SLCMJ Con	0.07	Slight
SLBJ Ecc – SLBJ Con	< 0.01	Slight
SLCMJ Ecc – SLBJ Con	0.21	Fair
SLBJ Ecc – SLCMJ Con	-0.25	Fair
Iso = isometric, SLCMJ = single leg countermovement jump, SLBJ = single leg broad jump, Ecc = eccentric, Con = concentric		

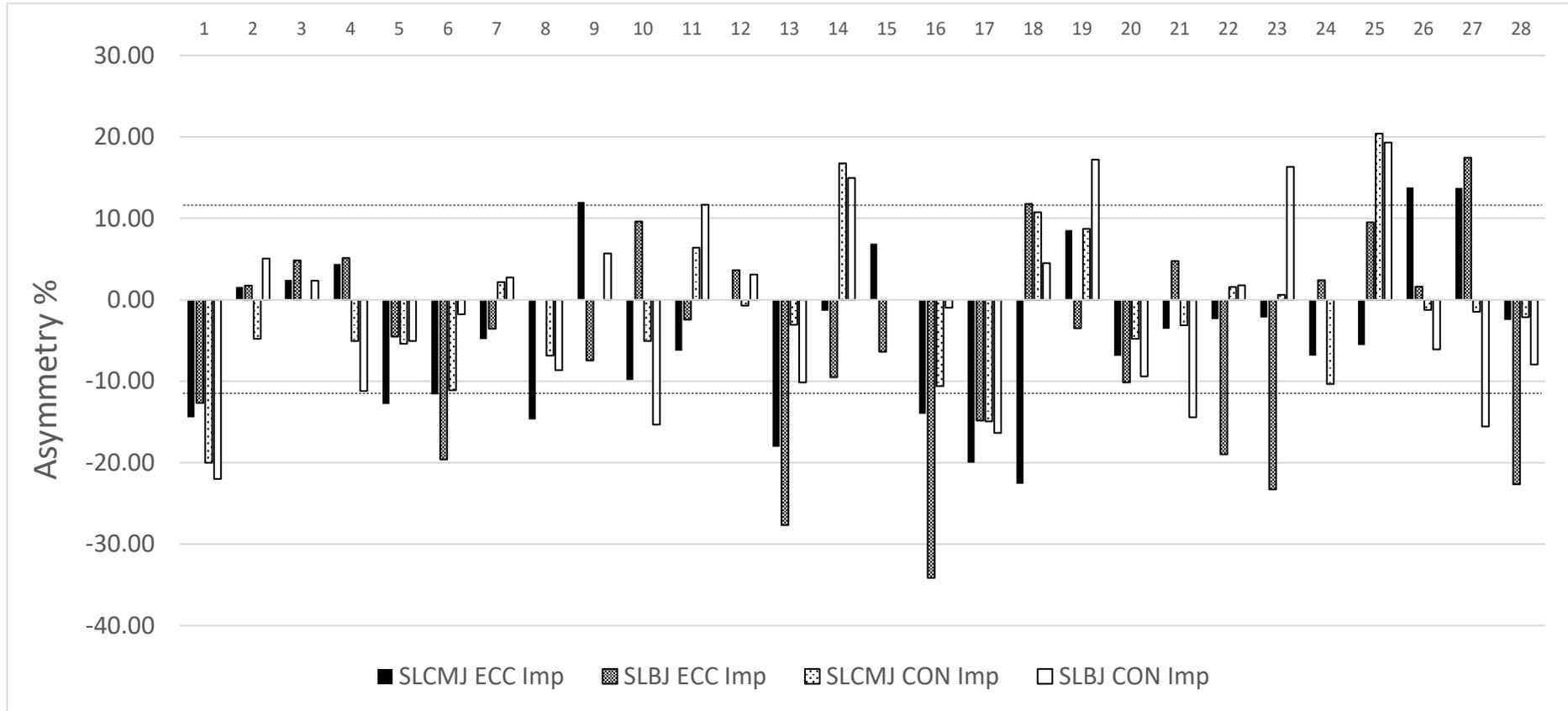
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400 Figure 1: Individual asymmetry data for peak force (PF) during the isometric squat (ISO Squat), single leg countermovement jump (SLCMJ),
 401 and single leg broad jump (SLBJ). Note: above the line indicates raw score is greater on the right limb and below the line indicates raw score is
 402 greater on the left limb. Dashed lines indicate largest coefficient of variation value for all PF measures.



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406 Figure 2: Individual asymmetry data for eccentric (ECC) and concentric (CON) impulse (Imp) during the single leg countermovement jump
 407 (SLCMJ) and single leg broad jump (SLBJ) tests. Note: above the line indicates raw score is greater on right limb and below the line indicates raw
 408 score is greater on left limb. Dashed lines indicate greatest coefficient of variation value for either eccentric or concentric impulse measures.